# A Two-Stage Group Sampling Plan Based on Truncated Life Tests for Exponentiated Half Logistic Distribution

G.S. Rao<sup>1\*</sup>, K. Rosaiah<sup>2</sup>, Ch.R. Naidu<sup>3</sup>

<sup>1</sup> Department of Mathematics and Statistics, The University of Dodoma Dodoma, PO. Box: 259, Tanzania

<sup>2</sup> Department of Statistics, Acharya Nagarjuna University Guntur – 522 007, India

> <sup>3</sup> Department of Statistics, Dilla University Dilla, PO. Box: 419, Ethiopia \*E-mail: gaddesrao@gmail.com

Received: 9 June 2020; revised: 26 June 2020; accepted: 27 June 2020; published online: 29 June 2020

Abstract: When the life time of a product follows exponentiated half logistic distribution we can recommend a two-stage group acceptance sampling plan for truncated life tests. The acceptance of the lot can be done at the first or second stage based on the number of failures in each group. In this paper, we obtained the number of groups essential for each of two stages for the underlying lifetime distribution so as to minimize the average sample number under the constraints of satisfying the producer's and consumer's risks simultaneously. Single-stage group sampling plans are also considered as special cases of the stated plan and compared with the proposed plan in terms of the average sample number and the operating characteristics.

Key words: exponentiated half logistic distribution, average sample number, consumer's risk, operating characteristics, producer's risk

## ANNOTATIONS

- $n_1$  sample size in stage 1,
- $n_2$  sample size in stage 2,
- $k_1$  number of testers in stage 1,
- $k_2$  number of testers in stage 2,
- m number of items assigned to the tester for testing at time  $t_0$ ,
- $c_1$  number of allowable defective items in the lot from stage 1,
- $c_2$  number of allowable defective items in the lot from stage 2,
- $P_{\rm a}^1$  probability of accepting the lot in stage 1,
- P<sub>a</sub><sup>2</sup> probability of accepting the lot in stage 2,
  P<sub>r</sub><sup>1</sup> probability of rejecting the lot in stage 1,

- $t_a^0$  targeted percentile at test termination time  $t_0$ ,
- p probability that a failure arises during termination time  $t_0$ ,
- $p_1$  failure probability corresponding to the consumer's risk,
- $p_2$  failure probability corresponding to the producer's risk,
- L(p) probability that the acceptance of the lot in the proposed two-stage group sampling plan,
- $L(p_1)$  lot acceptance probability corresponding to consumer's risk,
- $L(p_2)$  lot acceptance probability corresponding to producer's risk,
- $\alpha$  producer's risk,
- $\beta$  consumer's risk.

## I. INTRODUCTION

A commonly used method in quality control is acceptance sampling. A lot will be accepted if the number of failures does not exceed the acceptance number during the test time. The purpose is to conclude about the quality of lot or group of items from a sample. We can accept or reject the entire lot depending on what is found in the sample and the rejected lots can then be scuffled and re-used. Several authors suggested the acceptance sampling plans based on truncated life tests including [1–15].

Usually a tester can test a single product in these plans but in real-world the multiple items may be fitted at a time by the tester. Therefore, a group acceptance sampling plan must be designed and used and this plan is more required than the normal sampling plan in terms of the test time since the previous plan can test more items in a specified test time. In group acceptance sampling plans a simple decision rule is to accept the lot when the number of failures does not exceed the given number from each group. In sudden death testing the group acceptance sampling plans are very useful and these plans are suggested by several authors including [16–22].

In various applications in order to encounter the engineering design purpose the lowest percentile of lifetime is needed. In Refs. [23-26] and lately in [12] the acceptance sampling plans for the lifetime percentiles were developed. Almost all of the above discussed that the average lifetime of an item may not satisfy the constraints of an engineering design. A lot with smaller percentile value than the specified percentile value might be permitted by the acceptance sampling plans. Furthermore, many of the working life distributions are asymmetric. According to [27], the average life might not be sufficient to define the central tendency of the skewed distribution. Moreover, the exponentiated half logistic distribution is not symmetrical distribution; therefore, we need to study the two-stage grouped sampling plans based on percentiles. As far as we know, in the literature the twostage group sampling plans based on the exponentiated half logistic distribution using percentiles have not been established. Here the objective is to develop the two-stage group sampling plans for the exponentiated half logistic distribution percentiles.

The so far discussed acceptance sampling plans were established for a single sampling plan with respect to the size of a sample. In fact, a double sampling plan works better than a single sampling plan. Hence, Aslam et al. [28] truly said that there is a necessity of evolving a type of a double sampling plan for a life test with groups which are named as a two-stage group sampling plan. The size of the sample will certainly be further reduced by the two-stage group sampling plan. Recently, Aslam et al. [28] studied a time truncated two-stage group sampling plan for Weibull distribution. Aslam et al. [29] developed a two-stage group acceptance sampling plan based on truncated life tests for a general distribution. Rao [30] formulated a two-stage group sampling plan based on truncated life tests for an M-O extended exponential distribution. Rao et al. [31, 32] established a two-stage group sampling plan based on truncated life tests for an exponentiated Frechet distribution. Prasad et al. [33] studied a two stage group sampling plans based on truncated life tests for a Type II generalized log-logistic distribution. Rao et al. [34] developed a time-truncated twostage group acceptance sampling plan for an odds exponential log-logistic distribution. Sivakumar et al. [35] developed a two-stage group sampling plan based on truncated life tests for an odd generalized exponential log-logistic distribution.

The persistence of this article is to improve a two-stage group sampling plan for the truncated life tests if the lifetime of an item follows the exponentiated half logistic distribution established by Cordeiro et al. [36].

The probability density function (p.d.f.) and cumulative distribution function (c.d.f) of the exponentiated half logistic distribution, respectively, are given by:

$$f(t; \nu, \sigma) = \frac{2\nu \left(1 - e^{-t/\sigma}\right)^{\nu-1} e^{-t/\sigma}}{\sigma \left(1 - e^{-t/\sigma}\right)^{\nu+1}} , \ t \ge 0, \quad (1)$$

$$F(t; \nu, \sigma) = [F(t; \sigma)]^{\nu} = \left(\frac{1 - e^{-t/\sigma}}{1 + e^{-t/\sigma}}\right)^{\nu}, \ t \ge 0, \ (2)$$

where  $\sigma$  is the scale parameter and  $\nu$  is shape parameter.

The 100q-th percentile of the exponentiated half logistic distribution is given by

$$t_q = \sigma \eta_q$$
 where  $\eta_q = \left[ \ln \left( \frac{1 + q^{1/\nu}}{1 - q^{1/\nu}} \right) \right]$ . (3)

In this article, for truncated life tests the two-stage group sampling plan will be proposed if the life time of an item follows an exponentiated half logistic distribution. To minimize the mean sample number under the limitations of satisfying the producer's and consumer's risks we need to build the tables for obtaining the number of groups necessary in every step of the proposed plan simultaneously. The proposal of a two-stage plan is given in Sec. 2. The comparison with the single-stage group sampling plan is given in Sec. 3. The procedure is explained with industrial application in Sec. 4 and some inferences are given in Sec. 5.

## II. TWO-STAGE GROUP SAMPLING PLAN

Here we establish a two-stage group sampling plans for exponentiated half logistic distribution established. The authors suggested ensuring a two-stage group sampling plan such that the assumption of each group size "m" is prefixed for a tester.

**Stage 1.** At this stage select a random sample of size  $n_1$  from the lot, then each of  $k_1$  groups (or testers) will be assigned m items so that  $n_1 = mk_1$  and set them for testing at time  $t_0$ . If the number of failures in each group is  $c_1$  or less then the lot will be accepted. The test will be terminated and the lot will be rejected if the number of failures in any group is greater than  $c_2$  before time  $t_0$ . Otherwise go to stage 2.

**Stage 2.** At this stage select the random sample of size  $n_2$  from the lot, then each of  $k_2$  groups (or testers) will be assigned m items so that  $n_2 = mk_2$  and set them for testing at time  $t_0$ . If the number of failures in each group is  $c_1$  or less then the lot will be accepted. The test will be terminated and the lot will be rejected if the number of failures in any group is greater than  $c_1$  before time  $t_0$ .

The various sampling plans are related to the current group sampling plan. If m = 1 the current sampling plan is a case of a double sampling plan. The current two-stage group sampling plan will become a single-stage group sampling plan in case of  $c_1 = c_2$ . The number of groups from stage 1 and stage 2 are the design parameters of the current sampling plan. The acceptance numbers of  $c_1$  and  $c_2$ can be determined as well, but the average sample number (ASN) is minimizing under the sampling plan in case of  $c_1 = 0, c_2 = 1$ . The consumers usually prefer the lower acceptance numbers, so that the two-stage sampling plan with  $c_1 = 0, c_2 = 1$  is very useful. The probability of accepting the lot from stage 1 in the two-stage group sampling plan is given by

$$P_{\rm a}^{(1)} = \sum_{i=0}^{c_1} \left( \begin{array}{c} mk_1 \\ i \end{array} \right) \, p^i \, (1-p)^{mk_1-i} \,, \tag{4}$$

where p is the probability of an item which fails in a group in time  $t_0$ . It will be useful to obtain the test termination time  $t_0$  as a number of the stated percentile  $t_q^0$  such that  $t_0 = \delta_q^0 t_q^0$ for a constant  $\delta_q^0$ . Then the probability that a failure arises during the termination time  $t_0$  denoted as  $p = F(t_0; \nu, \sigma)$ , and is given by

$$p = \left[\frac{1 - e^{-\left(\frac{\eta_q \delta_q^0}{(t_q / t_q^0}\right)}}{1 + e^{-\left(\frac{\eta_q \delta_q^0}{(t_q / t_q^0)}\right)}}\right]^{\nu} .$$
 (5)

The probability that the lot will be rejected in stage 1 is given by

$$P_{\rm r}^{(1)} = 1 - \sum_{i=0}^{c_2} \left( \begin{array}{c} mk_1 \\ i \end{array} \right) \, p^i \, (1-p)^{mk_1-i} \,. \tag{6}$$

The probability that the lot will be accepted in stage 2 is given by

$$P_{\rm a}^{(2)} = \left[1 - \left(P_{\rm a}^{(1)} + P_{\rm r}^{(1)}\right)\right] \left[\sum_{i=0}^{c_1} \binom{mk_2}{i} p^i (1-p)^{mk_2-i}\right].$$
(7)

Then, the probability that the acceptance of the lot in the recommended two-stage group sampling plan is given by

$$L(p) = P_{\rm a}^{(1)} + P_{\rm a}^{(2)} \,. \tag{8}$$

But, the probability that the lot will be accepted if  $c_1 = 0$ ,  $c_2 = 1$ , is given by

$$L(p) = (1-p)^{mk_1} + mk_1p (1-p)^{mk_1-1} (1-p)^{mk_2}.$$
 (9)

The two-point approach of obtaining the proposal constraints if the quality level based on the percentile ratio  $t_q/t_q^0$  between the true percentile  $t_q$  and the target percentile  $t_q^0$  is to obtain the least number of groups,  $k_1$  and  $k_2$ , to satisfy the following two inequalities:

$$L\left(p/(t_q/t_q^0) = \delta_1\right) \le \beta, \qquad (10)$$

$$L\left(p/(t_q/t_q^0) = \delta_2\right) \ge 1 - \alpha \,, \tag{11}$$

where,  $\delta_1$  is the percentile ratio at the consumer's risk and  $\delta_2$  is the percentile ratio at the producer's risk. In this study, the ratio  $\delta_1$  is set as 1. Let  $p_1$  and  $p_2$  are the failure probabilities of corresponding to consumer's and producer's risks, respectively:

$$p_{1} = \left[\frac{1 - e^{-\eta_{q}\delta_{q}^{0}}}{1 + e^{-\eta_{q}\delta_{q}^{0}}}\right]^{\nu} \text{ and } p_{2} = \left[\frac{1 - e^{-\left(\frac{\eta_{q}\delta_{q}^{0}}{(t_{q}/t_{q}^{0})}\right)}}{1 + e^{-\left(\frac{\eta_{q}\delta_{q}^{0}}{(t_{q}/t_{q}^{0})}\right)}}\right]^{\nu}.$$
 (12)

Many solutions of the proposal constraints satisfying the Eqs. (10) and (11) may exist; therefore we must select to minimize the ASN for the two-stage group sampling plan. The ASN for the two-stage sampling plan can be determined as

ASN = 
$$mk_1 + mk_2 \left(1 - P_a^{(1)} - P_r^{(1)}\right)$$
, (13)

where  $P_{\rm a}^{(1)}$  and  $P_{\rm r}^{(1)}$  will be estimated at  $p = p_2$ . So, the proposal constraints for the recommended two-stage group sampling plan can be determined by the result from the following optimization:

Minimize 
$$ASN(p_2) = mk_1 + mk_2 \left(1 - P_a^{(1)} - P_r^{(1)}\right)$$
, (14a)  
subject to:

subject to:

$$L\left(p_{1}\right) \leq \beta,\tag{14b}$$

$$L\left(p_2\right) \ge 1 - \alpha,\tag{14c}$$

$$1 \le k_2 \le k_1 , \qquad (14d)$$

$$0 \le c_1 < c_2 , \qquad (14e)$$

$$k_1, k_2, c_1, c_2$$
 are integers. (14f)

If the number of groups in stage 2 is bigger than in stage 1, the constraint (14d) are not necessary.

Tab. 1. The minimum number of groups required in the two-stage sampling plan for EHLD with  $\nu = 2$  for  $25^{\text{th}}$  percentile. The cells with hyphens (–) indicate that parameters cannot be found to satisfy the conditions

B	4 /40	$\delta_q = 0.5 , \ m = 5$					$\delta_q = 1.0 , \ m = 5$				$\delta_q = 0$	0.5, m	= 10	$\delta_q = 1.0 , \ m = 10$			
ρ	$\iota_q/\iota_q$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$
	4	7	1	35.7	0.9849	2	1	10.8	0.9718	3	1	31.2	0.9855	1	1	11.6	0.9589
0.25	6	7	1	35.3	0.9968	2	1	10.4	0.9938	3	1	30.6	0.9969	1	1	10.8	0.9908
0.25	8	7	1	35.2	0.9990	2	2	10.5	0.9970	3	1	30.3	0.9990	1	1	10.5	0.9970
	10	7	1	35.1	0.9996	2	1	10.1	0.9992	3	1	30.2	0.9996	1	1	10.3	0.9987
0.10	4	10	1	50.9	0.9723	_	_	-	_	5	1	51.9	0.9681	1	1	11.6	0.9589
	6	10	1	50.5	0.9940	3	1	15.6	0.9887	5	1	50.9	0.9930	1	1	10.8	0.9908
	8	10	1	50.3	0.9980	3	1	15.3	0.9962	5	1	50.6	0.9977	1	1	10.5	0.9970
	10	10	1	50.2	0.9992	3	1	15.2	0.9984	5	1	50.4	0.9990	1	1	10.3	0.9987
	4	12	1	61.1	0.9624	_	_	-	_	6	1	62.1	0.9575	_	_	-	_
0.05	6	12	1	60.6	0.9917	3	1	15.6	0.9887	6	1	61.1	0.9906	2	1	21.4	0.9766
0.05	8	12	1	60.3	0.9973	3	1	15.3	0.9962	6	1	60.7	0.9969	2	1	20.9	0.9921
	10	12	1	60.2	0.9989	3	1	15.2	0.9984	6	1	60.4	0.9987	2	1	20.6	0.9966
	4	-	-	_	_	-	-	_	_	-	-	_	_	-	-	_	_
0.01	6	17	1	85.7	0.9845	4	1	20.7	0.9822	9	1	91.6	0.9812	2	1	21.4	0.9766
0.01	8	17	1	85.5	0.9948	4	1	20.4	0.9940	9	1	91.0	0.9937	2	1	20.9	0.9921
	10	17	1	85.3	0.9978	4	1	20.3	0.9975	9	1	90.6	0.9973	2	1	20.6	0.9966

Tab. 2. The minimum number of groups required in the two-stage sampling plan for EHLD with  $\nu = 3$  for  $25^{\text{th}}$  percentile. The cells with hyphens (–) indicate that parameters cannot be found to satisfy the conditions

в	+ /+0		$\delta_q =$	0.5, m	= 5	$\delta_q = 1.0 , \ m = 5$				$\delta_q = 0.5 , \ m = 10$					$\delta_q = 1.0 , \ m = 10$			
ρ	$\iota_q/\iota_q$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$	
	4	11	1	55.2	0.9989	2	1	10.2	0.9966	6	1	60.4	0.9986	1	1	10.5	0.9949	
0.25	6	11	1	55.0	0.9999	2	1	10.0	0.9997	6	1	60.1	0.9999	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	0.9995			
0.25	8	11	1	55.0	1.0000	2	1	10.0	0.9999	6	1	60.1	1.0000	1	1	10.1	0.9999	
	10	11	1	55.0	1.0000	2	1	10.0	1.0000	6	1	60.0	1.0000	1	1	10.0	1.0000	
	4	17	1	85.3	0.9976	3	1	15.4	0.9937	8	1	80.5	0.9976	1	1	10.5	0.9949	
0.10	6	17	1	85.1	0.9998	3	1	15.1	0.9994	8	1	80.1	0.9998	1	1	10.1	0.9995	
0.10	8	17	1	85.0	1.0000	3	1	15.1	0.9999	8	1	80.1	1.0000	1	1	10.1	0.9999	
	10	17	1	85.0	1.0000	3	1	15.0	1.0000	8	1	80.0	1.0000	1	1	10.0	1.0000	
	4	20	1	100.3	0.9968	3	1	15.4	0.9937	10	1	100.7	0.9965	2	1	10.0 21.1	0.9868	
0.05	6	20	1	100.1	0.9997	3	1	15.1	0.9994	10	1	100.2	0.9997	2	1	20.3	0.9987	
0.05	8	20	1	100.0	0.9999	3	1	15.1	0.9999	10	1	100.1	0.9999	2	1	20.1	0.9998	
	10	20	1	100.0	1.0000	3	1	15.0	1.0000	10	1	100.1	1.0000	2	1	20.1	0.9999	
	4	29	1	145.5	0.9936	4	1	20.5	0.9900	14	1	140.9	0.9936	2	1	21.1	0.9868	
0.01	6	29	1	145.2	0.9994	4	1	20.1	0.9990	14	1	140.3	0.9994	2	1	20.3	0.9987	
0.01	8	29	1	145.1	0.9999	4	1	20.1	0.9998	14	1	140.1	0.9999	2	1	20.1	0.9998	
	10	29	1	145.0	1.0000	4	1	20.0	1.0000	14	1	140.1	1.0000	2	1	20.1	0.9999	

Tab. 3. The minimum number of groups required in the two-stage sampling plan for EHLD with  $\nu = 1.57$  for  $25^{\text{th}}$  percentile. The cells with hyphens (–) indicate that parameters cannot be found to satisfy the conditions

β	$t / t^0$		$\delta_q =$	0.5, m	z = 5	$\delta_q = 1.0 , \ m = 5$				$\delta_q = 0.5 , \ m = 10$					$\delta_q = 1.0 , \ m = 10$			
ρ	$v_q / v_q$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$	$k_1$	$k_2$	ASN	$L(p_1)$	
	4	5	1	26.0	0.9612	2	1	10.7	0.9774	-	_	_	_	_	-	-	_	
0.25	6	5	1	25.6	0.9881	2	1	10.4	0.9904	3	1	31.4	0.9801	1	1	11.4	0.9670	
0.23	8	5	1	25.4	0.9950	2	1	10.3	0.9951	3	1	30.9	0.9915	1	1	10.9	0.9857	
	10	5	1	25.2	0.9975	-	_	_	_	3	1	30.7	0.9957	1	1	10.6	0.9926	
0.10	4	_	_	_	_	-	_	_	_	-	_	_	_	_	_	_	_	
	6	8	1	40.9	0.974	3	1	15.9	0.9595	4	1	41.7	0.9692	1	1	11.4	0.9670	
0.10	8	8	1	40.6	0.9888	3	1	15.6	0.9824	4	1	41.2	0.9867	1	1	10.9	0.9857	
	10	8	1	40.4	0.9943	3	1	15.5	0.9909	4	1	40.9	0.9932	1	1	10.6	0.9926	
	4	_	_	-	_	-	_	_	_	_	_	_	_	_	_	_	_	
0.05	6	10	1	51.0	0.9623	3	1	15.9	0.9595	5	1	52.1	0.9566	_	-	-	_	
0.05	8	10	1	50.7	0.9836	3	1	15.6	0.9824	5	1	51.4	0.981	2	1	21.7	0.9642	
	10	10	1	50.5	0.9915	3	1	15.5	0.9909	5	1	51.1	0.9902	2	1	21.2	0.9812	
	4	_	_	_	_	_	_	_	_	_	_	_	_	_	_	_	_	
0.01	6	-	-	-	_	-	_	-	_	-	_	-	_	_	-	-	_	
0.01	8	14	1	70.9	0.9706	4	1	20.8	0.9726	7	1	71.9	0.9673	2	1	21.7	0.9642	
	10	14	1	70.7	0.9847	4	1	20.6	0.9857	7	1	71.4	0.9829	2	1	21.2	0.9812	

Thus the plan parameters of the design  $k_1$  and  $k_2$  will be obtained for given  $\alpha$  and  $\beta$ , at the consumer's risk the percentile ratio  $\delta_1$  and at the producer's risk the percentile ratio  $\delta_2$  such that ASN( $p_2$ ) should be minimized and for given values of shape parameter  $\nu$ , test termination ratio  $\delta_a^0$ , and the number of testers m the inequalities (14b) and (14c) are satisfied simultaneously. The minimum number of groups required for the two-stage group sampling plan according to the values of percentile ratios  $\delta_2 = 4, 6, 8, 10$  and  $\delta_1 = 1$ when m = 5 and m = 10 at the four levels of the consumer's risks such as  $\beta = 0.25, 0.10, 0.05$  and 0.01 with known parameter  $\nu = 2,3$  and 1.57 are shown in Tabs. 1–3. From the tables we can also observe that the number of groups required decreases as the group size increases from m = 5to m = 10 when the other parameters remain the same and also the ASN slightly increases. As the group size increases the sample size  $(mk_1 \text{ or } mk_2)$  decreases, which shows that a larger group will be more economical. We have also found that the number of groups not inclined by the shape parameter.

## III. COMPARISONS WITH SINGLE-STAGE GROUP SAMPLING PLANS

We suggest a single-stage group sampling plan (with group size m) as a special case of a recommended two-stage group sampling plan and formulate the table for designing

the plan. As discussed earlier, this is the case when  $c_1 = c_2 = c$  in the two-stage group sampling plan. The probability that the acceptance of the lot under this plan is given by

$$P_{\rm a} = \sum_{i=0}^{c} \left( \begin{array}{c} mk \\ i \end{array} \right) \ p^{i} \ (1-p)^{mk-i} \ , \tag{15}$$

where k is the number of groups needed. Apparently, the ASN is attained by m times k. Similarly as in the two-stage group sampling plan, a table for determining the number of groups and acceptance number required can be prepared according to the value of the specified unreliability at a given consumer's risk.

The minimum number of groups required for the singlestage group sampling plan with c = 0 or c = 1 when m = 5and m = 10 according to the values of percentile ratios  $\delta_2 = 4, 6, 8, 10$  and  $\delta_1 = 1$  at the four levels of the consumer's risks such as  $\beta = 0.25, 0.10, 0.05$  and 0.01 with known parameter  $\nu = 2, 3$  and 1.57 are computed to save the space displayed in Tab. 4 for  $\nu = 3$  only.

In the sampling plans technique a plan is said to be better than the existing one if its ASN is lower than ASN of the existing plan. Here, when you compare the ASN values of the proposed sampling plan it shows less than the single sampling plan. Also when you see the OC curve of two-stage sampling plan, it is in between the OC curves of single sampling plans at c = 0 and at c = 1. We observe from Fig. 1 that the planned two-stage group sampling design performs

Tab. 4. The minimum number of groups required in the single-stage sampling plan for EHLD with  $\nu = 3.0$  for  $25^{\text{th}}$  percentile. The cells with hyphens (–) indicate that parameters cannot be found to satisfy the conditions

			$\nu =$	$= 3.0, \delta =$	= 0.5	, m =	5		ν =	= 3.0 , δ =	= 1.0	, m =	5
$\beta$	$t_q/t_q^0$	si	single with $c = 0$			single with $c = 1$			ngle wit	h c = 0	single with $c = 1$		
		k	ASN	OC	k	ASN	OC	k	ASN	OC	k	ASN	OC
	4	7	35	0.9730	12	60	0.9989	1	5	0.97	2	10	0.9984
0.25	6	7	35	0.9918	12	60	0.9999	1	5	0.9907	2	10	0.9998
0.23	8	7	35	0.9965	12	60	1.0000	1	5	0.996	2	10	1.0000
	10	7	35	0.9982	12	60	1.0000	1	5	0.9980	2	10	1.0000
	4	11	55	0.9575	18	90	0.9976	_	-	-	3	15	0.9962
0.10	6	11	55	0.9872	18	90	0.9998	2	10	0.9816	3	15	0.9996
0.10	8	11	55	0.9946	18	90	1.0000	2	10	0.9921	3	15	0.9999
	10	11	55	0.9972	18	90	1.0000	2	10	0.9960	3	$\begin{array}{c} m = 1\\ m $	1.0000
	4	_	_	_	21	105	0.9968	_	-	-	4	20	0.9933
0.05	6	14	70	0.9837	21	105	0.9997	3	15	0.9725	4	20	0.9994
0.05	8	14	70	0.9931	21	105	0.9999	3	15	0.9882	4	20	0.9999
	10	14	70	0.9964	21	105	1.0000	3	15	0.9939	4	20	1.0000
	4	-	_	-	30	150	0.9936	_	_	-	5	25	0.9897
0.01	6	21	105	0.9756	30	150	0.9994	4	20	0.964	5	25	0.9990
0.01	8	21	105	0.9896	30	150	0.9999	4	20	0.9843	5	25	0.9998
	10	21	105	0.9947	30	150	1.0000	4	20	0.9919	5	25	1.0000
			ν	v = 3.0,	$\delta = 0$	0.5, m	= 10		ν	r = 3.0,	$\delta = 1$	1.0, m =	= 10
	4	4	40	0.9689	6	60	0.9989	-	_	-	1	10	0.9984
0.25	6	4	40	0.9906	6	60	0.9999	1	10	0.9816	1	10	0.9998
	8	4	40	0.9960	6	60	1.0000	1	10	0.9921	1	10	1.0000
	10	4	40	0.9980	6	60	1.0000	1	10	0.9960	1	10	1.0000
	4	6	60	0.9537	9	90	0.9976	-	-	-	2	20	0.9933
0.10	6	6	60	0.9860	9	90	0.9998	1	10	0.9816	2	20	0.9994
	8	6	60	0.9941	9	90	1.0000	1	10	0.9921	2	20	0.9999
	10	6	60	0.9970	9	90	1.0000	1	10	0.9960	2	20	1.0000
	4	-	-	-	11	110	0.9965	-	-	-	2	20	0.9933
0.05	6	7	70	0.9837	11	110	0.9997	2	20	0.9635	2	20	0.9994
	8	7	70	0.9931	11	110	0.9999	2	20	0.9843	2	20	0.9999
	10	7	70	0.9964	11	110	1.0000	2	20	0.9919	2	20	1.0000
	4	-	-	-	15	150	0.9936	-	-	-	3	30	0.9853
0.01	6	11	110	0.9745	15	150	0.9994	2	20	0.9635	3	30	0.9985
-	8	11	110	0.9891	15	150	0.9999	2	20	0.9843	3	30	0.9997
	10	11	110	0.9944	15	150	1.0000	2	20	0.9919	3	30	0.9999

			$\nu =$	$1.57, \delta$	= 0.5	5, m =	5		$\nu =$	$1.57, \delta$	= 1.0	0, m =	5
$\beta$	$t_q/t_q^0$	si	ngle wit	h c = 0		single w	ith $c = 1$	si	ngle wit	th $c = 0$		single wi	ith $c = 1$
		k	ASN	OC	k	ASN	OC	k	ASN	OC	k	$\begin{array}{c} 0 \ , \ m = \\ \text{single with } \\ \hline \text{ASN} \\ \hline 10 \\ 10 \\ 10 \\ 10 \\ 10 \\ 10 \\ \hline 10 \\ 10 \\$	OC
	4	-	_	_	6	30	0.9604	-	_	_	2	10	0.9632
0.25	6	-	_	-	6	30	0.9878	-	-	-	2	10	0.9888
0.25	8	_	_	-	6	30	0.9949	_	_	_	2	10	0.9953
	10	3	15	0.9631	6	30	0.9974	1	5	0.9635	2	10	0.9976
	4	-	_	-	-	_	_	-	_	-	-	_	_
0.10	6	_	_	-	9	45	0.9738	_	_	_	3	15	0.9752
0.10	8	_	_	-	9	45	0.9887	_	_	_	3	15	0.9894
	10	-	_	_	9	45	0.9942	-	_	-	3	15	0.9946
	4	_	_	-				-	-	-			
0.05	6	-	_	-	11	55	0.9620	_	_	-	4	20	0.9575
0.05	8	_	_	-	11	55	0.9835	_	_	_	4	20	0.9815
	10	-	_	-	11	55	0.9915	_	_	-	4	20	0.9904
	4	_	_	-	_	-	_	_	-	_	_	-	_
0.01	6	-	_	_	-	_	_	-	_	-	-	_	_
0.01	8	-	_	-	5	25	0.9717	-	_	_	5	50	0.9717
	10	-	_	_	5	25	0.9853	-	_	-	5	50	0.9853
			ν	= 1.57,	$\delta =$	0.5, m	= 10		ν	= 1.57,	$\delta =$	1.0, m	= 10
	4	-	_	-	3	30	0.9604	-	_	-	1	10	0.9632
0.25	6	-	_	-	3	30	0.9878	-	_	_	1	10	0.9888
0.25	8	-	_	-	3	30	0.9949	-	_	-	1	10	0.9953
	10	2	20	0.9511	3	30	0.9974	_	_	_	1	10	0.9976
	4	-	_	_	-	-	_	-	_	_	-	-	_
0.10	6	-	_	-	5	50	0.9681	-	_	_	2	20	0.9575
0.10	8	-	_	-	5	50	0.9862	-	_	_	2	20	0.9815
	10	-	_	_	5	50	0.9929	-	_	-	2	$\begin{array}{c} m = \\ single wi \\ ASN \\ \hline 10 \\ 10 \\ 10 \\ 10 \\ 10 \\ \hline 10 \\ 10 \\$	0.9904
	4	-	_	_				-	_	_			
0.05	6	-	-	-	6	60	0.9556	_	_	-	2	20	0.9575
0.05	8	-	-	-	6	60	0.9805	_	_	-	2	20	0.9815
	10	-	-	-	6	60	0.9899	_	_	-	2	20	0.9904
	4	-	_	_	-	_	_	_	_	_	_	_	_
0.01	6	-	-	-	_	-	_	_	-	-	_	-	_
0.01	8	_	_	_	8	80	0.9668	_	_	_	3	30	0.9604

8

\_

80

0.9826 –

\_

10

\_

\_

Tab. 5. The minimum number of groups required in the single-stage sampling plan for EHLD with  $\nu = 1.57$  for  $25^{\text{th}}$  percentile. The cells with hyphens (–) indicate that parameters cannot be found to satisfy the conditions

0.9791

3

\_

30



Fig. 1. The OC curves for ASN for different sampling plans

better than the single-stage group sampling plan in terms of the ASN and the OC values (see Tab. 2 with Tab. 4). While comparing the OC curves from Fig.1 at the chosen plan values of  $m = 10, \nu = 3, \beta = 0.25, t_q/t_q^0 = 4, \delta = 0.5, t_q = 0.25 (25^{\rm th} \text{ percentile})$  the  $k_1 = 6$  and  $k_2 = 1$  from Tab. 2 represented by the middle curve for ASN for a two-stage sampling plan with  $c_1 = 0$  and  $c_2 = 1$ , and the k = 4, k = 6 from Tab. 4 represented by lower and upper curves for the ASN for a single stage sampling plan with c = 0 and c = 1, respectively. We also observed that the ASN for the two-stage group sampling plan with  $c_1 = 0, c_2 = 1$  is much smaller than that of the single-stage group sampling plan with c = 1.

#### **IV. INDUSTRIAL APPLICATIONS**

At this juncture we use the real data set to demonstrate the methodology of a two-stage group sampling plan for ex-

(a) Empirical and Fitted PDFs

ponentiated half logistic distribution. The real dataset consists of 76 observations of exact times of failures, and this data set was previously used by Abdul-Moniem and Seham [37] and Andrews and Herzberg [38]. For quick reference the data set is reported here, the times of failures (hours): 0.0251, 0.0886, 0.0891, 0.2501, 0.3113, 0.3451, 0.4763, 0.5650, 0.5671, 0.6566, 0.6748, 0.6751, 0.6753, 0.7696, 0.8375, 0.8391, 0.8425, 0.8645, 0.8851, 0.9113, 0.9120, 0.9836, 1.0483, 1.0596, 1.0773, 1.1733, 1.2570, 1.2766, 1.2985, 1.3211, 1.3503, 1.3551, 1.4595, 1.4880, 1.5728, 1.5733, 1.7083, 1.7263, 1.7460, 1.7630, 1.7746, 1.8275, 1.8375, 1.8503, 1.8808, 1.8878, 1.8881, 1.9316, 1.9558, 2.0048, 2.0408, 2.0903, 2.1093, 2.1330, 2.2100, 2.2460, 2.2878, 2.3203, 2.3470, 2.3513, 2.4951, 2.5260, 2.9911, 3.0256, 3.2678, 3.4045, 3.4846, 3.7433, 3.7455, 3.9143, 4.8073, 5.4005, 5.4435, 5.5295, 6.5541, 9.0960. This data set is used to demonstrate the goodness of fit for the exponentiated half logistic distribution and hence methodology of the two-stage group sampling plans. Fig. 2 depicts the estimated density and QQ plot to emphasize the goodness of EHLD for times of failures data set. The maximum likelihood estimates of the parameter of EHLD for the exact times of failures is  $\nu = 1.57$ , the Kolmogorov-Smirnov test for the maximum distance between the data and the fitted of the EHLD is 0.08903 with *p*-value is 0.553. Therefore, the two-parameter EHLD provides reasonable fits for lifetimes of items.

Suppose that an experimenter would like to establish the true unknown  $25^{\text{th}}$  percentile lifetime for the product for at least 0.1 hours and the experiment will be stopped after 0.2 hours, which leads to  $t_q/t_{q_0} = 2$ . Assume that times of failures follow EHLD with  $\nu = 1.57$ , in the laboratory the experimenter has facility to install five items on a group. Suppose that  $\beta = 0.10$  and  $t_q/t_{q_0} = 2$  with  $\alpha = 0.05$  for this experiment from Tab. 3, the two-stage acceptance



Fig. 2. (a) Estimated density and (b) Q-Q plot for times of failures data

sampling plan parameters at m = 5 and  $\delta = 0.5$  are  $k_1 = 5, k_2 = 1, c_1 = 0$  and  $c_2 = 1$ . The two-stage acceptance sampling plan is implemented as follows: randomly select 25 items and distribute 5 items into each of 5 testers and accept the product if there is no failure from each tester in 0.05 hours and reject the product if there is more than 1 failure from any tester before 0.05 hours. If one failure is observed from any tester then go to the second stage, then randomly select another 5 items from the lot and distribute 5 items from the two-stage testing within 0.05 hours for each stage is less than one then the lot is accepted; otherwise, the lot is rejected. In this example, only one failure before 0.05 hours hence lot is accepted. The probability of acceptance for this plan is 0.9612 and ASN is 26.0.

#### V. CONCLUSION

In this study, a two-stage group acceptance sampling plan is developed when the lifetime of a product follows exponentiated half logistic distribution percentiles with a known shape parameter. The plan parameters like the number of groups in each stage were determined so as to minimize the ASN subject to satisfy the consumer's and the producer's risks simultaneously under several conditions. Tables for the plan parameters were constructed under various combinations such as the known shape parameter, group sizes, consumer's and the producer's risks, and so on. An industrial example has been presented to illustrate the applications of the proposed two-stage group sampling plan. We made the comparison between the proposed two-stage group acceptance sampling plan and single-stage acceptance sampling plan. We observed from tables that the numbers of groups required decrease as the group size increases from 5 to 10 when other parameters remain the same and also the ASN increases marginally. The sample size also decreases as the group size increases, which indicates that a larger group size may be more economical. Finally, it should be mentioned from tables that the proposed two-stage group acceptance sampling plan performs better in terms of the average sample number and the operating characteristics than a single-stage group acceptance sampling plan.

#### References

- [1] B. Epstein, *Truncated life tests in the exponential case*, The Annals of Mathematical Statistics **25**(3), 555–564 (1954).
- [2] M. Sobel, J.A. Tischendrof, Acceptance sampling with sew life test objective, Proceedings of Fifth National Symposium on Reliability and Quality Control, 108–118, Philadelphia (1959).
- [3] S.S. Gupta, P.A. Groll, *Gamma distribution in acceptance sampling based on life tests*, Journal of the American Statistical Association 56, 942–970 (1961).

- [4] H. Goode, J. Kao, Sampling plans based on the Weibull distribution, Proceeding of the Seventh National Symposium on Reliability and Quality Control, 24–40 (1961).
- [5] S. Gupta, *Life test sampling plans for normal and lognormal distributions*, Technometrics **4**(2), 151–175 (1962).
- [6] F. Fertig, N. Mann, Life-test sampling plans for two-parameter Weibull populations, Technometrics 22(2), 165–177 (1980).
- [7] R.R.L. Kantam, K. Rosaiah, *Half logistic distribution in acceptance sampling based on life tests*, IAPQR Transactions 23(2), 117–125 (1998).
- [8] R.R.L. Kantam, K. Rosaiah, G.S. Rao, Acceptance sampling based on life tests: Log- logistic models, Journal of Applied Statistics 28(1), 121–128 (2001).
- [9] A. Baklizi, Acceptance sampling based on truncated life tests in the Pareto distribution of the second kind, Advances and Applications in Statistics 3(1), 33–48 (2003).
- [10] A. Baklizi, A. EI Masri, Acceptance sampling based on truncated life tests in the Birnbaum-Saunders model, Risk Analysis 24(6), 1453–1457 (2004).
- [11] T.-R. Tsai, S.-J. Wu, Acceptance sampling based on truncated life tests for generalized Rayleigh distribution, Journal of Applied Statistics 33(6), 595–600 (2006).
- [12] N. Balakrishnan, V. Leiva, J. Lopez, Acceptance sampling plans from truncated life tests based on the generalized Birnbaum-Saunders distribution, Communication in Statistics-Simulation and Computation 36, 643–656 (2007).
- [13] M. Aslam, Double acceptance sampling based on truncated life tests in Rayleigh distribution, European Journal of Scientific Research 7(4), 605–610 (2007).
- [14] G.S. Rao, M.E. Ghitany, R.R.L. Kantam, Acceptance sampling plans for Marshall-Olkin extended Lomax distribution, International Journal of Applied Mathematics 21(2), 315– 325 (2008).
- [15] A.I. Al-Omari, Acceptance sampling plans based on truncated life tests for sushila distribution, Journal of mathematical and fundamental sciences 50(1), 72–83 (2018).
- [16] G.S. Rao, A group acceptance sampling plans based on truncated life tests for generalized exponential distribution, Economic Quality Control 24(1), 75–85 (2009).
- [17] G.S. Rao, A group acceptance sampling plans based on truncated life tests for Marshall-Olkin extended Lomax distribution, Electronic Journal of Applied Statistical Analysis 3(1), 18–27 (2010).
- [18] M. Aslam, C.-H. Jun, A group acceptance sampling plan for truncated life test having Weibull distribution, Journal of Applied Statistics 39, 1021–1027 (2009).
- [19] M. Aslam, C.-H. Jun, A group acceptance sampling plans for truncated life tests based on the inverse Rayleigh and log-logistic distributions, Pakistan Journal of Statistics 25, 107–119 (2009).
- [20] C.-H. Jun, S. Balamurali, S.-H. Lee, Variables sampling plans for Weibull distributed lifetimes under sudden death testing, IEEE Transactions on Reliability 55(1), 53–58 (2006).
- [21] J.-W. Wu, T.-R. Tsai, L.-Y. Ouyang, *Limited failure-censored life test for the Weibull distribution*, IEEE Transactions on Reliability 50(1), 107–111 (2001).
- [22] F. Pascual, W. Meeker, *The modified sudden death test: planning life tests with a limited number of test positions*, Journal of Testing and Evaluation 26(5), 434–443 (1998).
- [23] G.S. Rao, Acceptance sampling plans from truncated life tests based on the Marshall-Olkin extended exponential distribution for percentiles, Brazilian Journal of Probability and Statistics 27(2), 117–132 (2013).

- [24] G.S. Rao, R.R.L. Kantam, Acceptance sampling plans from truncated life tests based on the log-logistic distribution for percentiles, Economic Quality Control 25(2), 153–167 (2010).
- [25] Y. Lio, T.-R. Tsai, S.-J. Wu, Acceptance sampling plan based on the truncated life test in the Birnbaum Saunders distribution for percentiles, Communications in Statistics-Simulation and Computation 39, 119–136 (2009).
- [26] Y.L. Lio, T.-R. Tsai, S.-J. Wu, Acceptance sampling plans from truncated life tests based on the Burr type XII percentiles, Journal of the Chinese Institute of Industrial Engineers 27(4), 270–280 (2010).
- [27] A.W. Marshall, I. Olkin, Life Distributions-Structure of Nonparametric, Semiparametric, and Parametric Families, Springer Series in Statistics (2007).
- [28] M. Aslam, C.-H. Jun, M. Rasool, M. Ahmad, A time truncated two-stage group sampling plan for Weibull distribution, Communications of the Korean Statistical Society 17, 89–98 (2010).
- [29] M. Aslam, C.-H. Jun, M. Ahmad, A two-stage group sampling plan based on truncated life tests for a general distribution, Journal of Statistical Computation and Simulation 81(12), 1927–1938 (2011).
- [30] G.S. Rao, A Two-stage group sampling plan based on truncated life tests for a M-O extended exponential distribution, International Journal of Quality Engineering & Technology 3(4), 319–331 (2013).
- [31] G.S. Rao, K. Rosaiah, M. Sridhar babu, D.C.U. Sivakumar, A two-stage group acceptance sampling plans based on trun-

*cated life tests for exponentiated frechet distribution*, European Scientific Journal **10**(33), 145–160 (2014).

- [32] G.S. Rao, C.N. Ramesh, Acceptance sampling plans for percentiles based on the exponentiated half logistic distribution, Applications and Applied Mathematics: An International Journal 9(1), 39–53 (2014).
- [33] S.V. Prasad, K. Rosaiah, G.S. Rao, A two stage group sampling plans based on truncated life tests for Type II generalized log-logistic distribution, International Journal of Scientific Research in Mathematical and Statistical Sciences 5(6), 228–243 (2018).
- [34] G.S. Rao, K. Kalyani, K. Rosaiah, D.C.U. Sivakumar, A timetruncated two-stage group acceptance sampling plan for odds exponential log-logistic distribution, Life Cycle Reliability and Safety Engineering 8(4), 337–345 (2019).
- [35] D.C.U. Sivakumar, G.S. Rao, K. Rosaiah, K. Kalyani, A twostage group acceptance sampling plans based on truncated life tests for an odd generalized exponential log-logistic distribution, Current journal of applied science and technology 35(4), 1–14 (2019).
- [36] G. Cordeiro, M. Alizadeh, E. Ortega, *The exponentiated half logistic family of distributions: properties and applications*, Journal of Probability and Statistics **2014**, 1–21 (2014).
- [37] I. Abdul-Moniem, M. Seham, *Transmuted Gompertz distribution*, Computational and Applied Mathematics 1(3), 88–96 (2015).
- [38] D.F. Andrews, A.M. Herzberg, Data: A Collection of Problems from Many Fields for the Student and Research Worker, Springer Series in Statistics (2012).



**Gadde Srinivasa Rao** received his MSc in Statistics (1988), MPhil in Statistics (1994) and PhD in Statistics (2002) from The Acharya Nagarjuna University, Guntur, India. He is presently working as Professor of Statistics at the Department of Mathematics and Statistics, The University of Dodoma, Tanzania. He boasts more than 120 publications in different peer-reviewed journals in national and international well reputed journals including the Journal of Applied Statistics, Quality and Reliability Engineering International, International Journal of Advanced Manufacturer Technology, Communications in Statistics Theory and methods, Communications in Statistics Simulation and Computation, Journal of Testing and Evaluation, Arabian Journal for Science and Engineering, International Journal of Quality & Reliability Management, Economic Quality Control and Journal of Statistical Computation and Simulation. He is the reviewer for various reputed international journals. His research interests include statistical inference, statistical process control, applied statistics, acceptance sampling plans and reliability estimation.



**Kanaparthi Rosaiah** has completed his MSc in Statistics in 1982 and got PhD in Statistics in 1990 from The Acharya Nagarjuan University (A.N.U.), Guntur, India. As Lecturer he joined the Dept. of Statistics, A.N.U. in 1985 and promoted as Professor in 2002. He has published 74 research articles in peer-reviewed national and international journals. 4 PhD and 16 MPhil degrees have been awarded to the research students under his guidance. At present 5 PhD and 1 MPhil students are pursuing under his guidance. His research interests include statistical inference, statistical quality control and reliability estimation.



**Chukka Ramesh Naidu** is Lecturer in Statistics in the Department of Statistics, Dilla University, Dilla, Ethiopia. He has been working as teaching faculty member since 2003, in the same institution since 2010. He received his MSc in Statistics in 2003, and he is doing his PhD in Statistics as NRI research scholar from The Acharya Nagarjuna University, Guntur, A.P., INDIA. His research area of interest is in life testing and reliability estimation in statistical quality control and applications of exponentiated half logistic distribution, and he has written total 9 research papers, 5 published and 4 communicated in various reputed national and international journals.